

Classification of Random Processes

- Discrete Random Sequence
- Discrete Random Process
- Continuous Random Sequence
- Continuous Random Process
- Deterministic Random Process
- Non Deterministic Random Process

1) Discrete Random Sequence

If both T and S are discrete, the random process is called Discrete Random Sequence.

Eg:- X_n represents the outcome of n th toss of a fair dice then $\{X_n, n \geq 1\}$ is a discrete random sequence, since $T = \{1, 2, 3, \dots\}$ & $S = \{1, 2, 3, 4, 5, 6\}$

2) Discrete Random Process

If T is continuous and S is discrete, the random process is called Discrete Random Process.

Eg:- X_t represents number of telephone calls received in the interval $(0, t)$. Then $\{X(t)\}$ is a discrete random process since $S = \{0, 1, 2, 3, \dots\}$

3) Continuous Random Sequence

If T is discrete and S is continuous, the random process is called a continuous Random Sequence.

Eg: X_n represents the temperature at the end of n th hour of the day then $\{X_n, 1 \leq n \leq 24\}$ is a Continuous Random Sequence since temperature can take any value in an interval and hence continuous.

4) Continuous Random process

If both T and S are continuous, the random process is called Continuous Random process.

Eg: If $X(t)$ represents maximum temperature at a place in the interval $(0, t)$, $\{X(t)\}$ is a continuous Random process.

Note :- Word 'Discrete' or 'Continuous' is used to refer to the nature of 'S'.

Word 'Sequence' or 'process' is used to refer to the nature of T .

5) Deterministic Random Process

If the future value of any sample function can be predicted from the knowledge of the past value, then the process is called as Deterministic Random process.

6) Non-Deterministic Random Process

If the future value of any sample function cannot be predicted from the knowledge of past value, then the process is called Non-Deterministic Random process.

Ridhi Sethi

Average values of Random Processes or Statistical averages of Random processes

(1) Mean

Mean of the process $\{X(t)\}$ is the expected value of a typical member $X(t)$ of the process.

$$\mu(t) = E\{X(t)\}$$

$$E\{X(t)\} = \int_{-\infty}^{\infty} x f(x, t) dx$$

(2) Auto - Correlation

Autocorrelation of the process $\{X(t)\}$ denoted by $R_{xx}(t_1, t_2)$ or $R_x(t_1, t_2)$ or $R(t_1, t_2)$ is the expected value of the product of any two members $X(t_1)$ and $X(t_2)$ of the process.

$$R(t_1, t_2) = E\{X(t_1) \times X(t_2)\}$$

(3) Autocovariance

Autocovariance of the process $\{X(t)\}$, denoted by $C_{xx}(t_1, t_2)$ or $C_x(t_1, t_2)$ or $C(t_1, t_2)$ is defined as

$$C(t_1, t_2) = E\left[\{X(t_1) - \mu(t_1)\} \{X(t_2) - \mu(t_2)\}\right]$$

$$= R(t_1, t_2) - \mu(t_1) \times \mu(t_2)$$

$$= E\{X(t_1) \times X(t_2)\} - \mu(t_1) \times \mu(t_2)$$

(4) Correlation Coefficient

Correlation Coefficient of the process denoted by $\rho_{xx}(t_1, t_2)$ or $\rho(t_1, t_2)$ is defined as

$$\rho(t_1, t_2) = \frac{C(t_1, t_2)}{\sqrt{C(t_1, t_1) C(t_2, t_2)}}$$

Where $C(t_1, t_1)$ is the variance of $X(t_1)$

* When we deal with two ^{or more} random processes, we can use joint distribution functions or averages to describe the relationship between them.

(5) Cross - Correlation of 2 processes

Cross - correlation of 2 processes $\{X(t)\}$ and $\{Y(t)\}$ is defined as

$$R_{xy}(t_1, t_2) = E\{X(t_1) \times Y(t_2)\}$$

(6) Cross - Covariance

Cross - covariance of 2 processes $\{X(t)\}$ and $\{Y(t)\}$ is defined as

$$C_{xy}(t_1, t_2) = R_{xy}(t_1, t_2) - \mu_x(t_1) \times \mu_y(t_2)$$

(7) Cross - Correlation Coefficient

Cross correlation coefficient of 2 processes $\{X(t)\}$ and $\{Y(t)\}$ is defined as

$$\rho_{xy}(t_1, t_2) = \frac{C_{xy}(t_1, t_2)}{\sqrt{C_{xx}(t_1, t_1) \times C_{yy}(t_2, t_2)}}$$

STATIONARITY | Stationary Random Process

A random process is called a strongly stationary process or strict sense stationary process (SSS) if all its finite dimensional distributions are invariant under translation of time period (shift in time eg:- $X(t)$ and $X(t+T)$ possesses same properties i.e. if the joint distributions (Joint density) of $X(t_1), X(t_2), \dots, X(t_n)$ is the same as that of $X(t_1+h), X(t_2+h), X(t_3+h), \dots, X(t_n+h)$ for all t_1, t_2, \dots, t_n and $h(>0)$ and for all $n \geq 1$, then the random process is called SSS process.

First order stationary Random process

A random process is said to be first order stationary if its first order density function is invariant under shift in time (translation in time) i.e. densities of $X(t)$ and $X(t+h)$ are same

$$f(x, t) = f(x, t+h)$$

This is possible if $f(x, t)$ is independent of t
 $\therefore E\{X(t)\}$ is also independent of t
 $\mu = \text{constant}$

Second order Stationary process

A random process is said to be second order stationary if the joint probability density functions (pdf) of $\{X(t_1)\}$ and $\{X(t_2)\}$ does not change with shift (translation) in time.

Joint pdf of $\{X(t_1), X(t_2)\}$ is same as $\{X(t_1+h), X(t_2+h)\}$

$$f(x_1, x_2, t_1, t_2) = f(x_1, x_2, t_1+h, t_2+h)$$

Ridhi Sethi

This is possible only if $f(x_1, x_2, t_1, t_2)$ is a function of $t = t_1 - t_2$

WIDE SENSE STATIONARITY (WSS)

A random process $\{X(t)\}$ with finite first order and second order moments is called a WEAKLY STATIONARY PROCESS OR COVARIANCE STATIONARY PROCESS OR WIDE SENSE STATIONARY PROCESS (WSS) if

(i) Its mean is constant

$$E\{X(t)\} = \mu = \text{constant}$$

(ii) Autocorrelation depends only on time difference.

$$E\{X(t) \times X(t - \tau)\} = R(\tau)$$

★ SSS process with finite first and second order moments is a WSS while a WSS process need not to be a SSS process.

★ A random process that is not stationary in any sense is called an evolutionary process.

★ Two random processes $\{X(t)\}$ and $\{Y(t)\}$ are said to be Jointly Stationary in the wide sense, if each process is individually a WSS process and $R_{xy}(t_1, t_2)$ is a function of $(t_1 - t_2)$ only.

Q:- The process $\{X(t)\}$ whose probability distribution under certain conditions is given by

$$P\{\{X(t)\} = n\} = \frac{(at)^{n-1}}{(1+at)^{n+1}}, \quad n=1, 2, 3, \dots$$

$$= \frac{at}{1+at}$$

Show that it is not stationary.

Sol:- For a process to be stationary:-

- (i) Mean $E[\{X(t)\}]$ is a constant -
- (ii) $E[\{X^2(t)\}]$ is constant -
- (iii) $R_{xx}(t_1, t_2)$ is a function of $(t_1 - t_2)$
- (iv) $\text{Var}\{X(t)\}$ is a constant.

The probability distribution of $X(t)$ is

$$X(t) = n \quad 0 \quad 1 \quad 2 \quad 3 \quad \dots$$

$$p_n \quad \frac{at}{1+at} \quad \frac{1}{(1+at)^2} \quad \frac{at}{(1+at)^3} \quad \frac{(at)^2}{(1+at)^4}$$

$$E\{X(t)\} = \sum_{n=0}^{\infty} n p_n$$

$$E\{X(t)\} = \frac{1}{(1+at)^2} + \frac{2at}{(1+at)^3} + \frac{3(at)^2}{(1+at)^4} + \dots \infty$$

$$= \frac{1}{(1+at)^2} \left[1 + \frac{2at}{1+at} + \frac{3(at)^2}{(1+at)^2} + \dots \right]$$

$$\text{Put } \frac{at}{1+at} = x$$

$$= \frac{1}{(1+at)^2} \left[1 + 2x + 3x^2 + \dots \right]$$

$$= \frac{1}{(1+at)^2} (1+x)^{-2}$$

$$= \frac{1}{(1+at)^2} \left[1 - \frac{at}{1+at} \right]^{-2} = \frac{1}{(1+at)^2} \left(\frac{1+at-at}{1+at} \right)^{-2}$$

$$\boxed{E\{X(t)\} = 1}$$

$$E\{X^2(t)\} = \sum_{n=0}^{\infty} n^2 p_n = 0 + \sum_{n=1}^{\infty} n^2 p_n$$

$$= \sum_{n=1}^{\infty} n(n+1) p_n - \sum_{n=1}^{\infty} n p_n$$

$$= \sum_{n=1}^{\infty} n(n+1) \frac{(at)^{n-1}}{(1+at)^{n+1}} - 1$$

$$\therefore \sum_{n=1}^{\infty} n p_n = \text{Mean}$$

$$E\{X^2(t)\} = \left[\frac{2}{(1+at)^2} + \frac{2(3)at}{(1+at)^3} + \frac{3(4)(at)^2}{(1+at)^4} + \dots \right] - 1$$

$$= \frac{2}{(1+at)^2} \left[1 + \frac{3at}{1+at} + \frac{6(at)^2}{(1+at)^2} + \dots \right] - 1$$

$$= \frac{2}{(1+at)^2} \left[1 + 3x + 6x^2 + \dots \right]$$

$$= \frac{2}{(1+at)^2} (1-x)^{-3}$$

where $x = \frac{at}{1+at}$

$$= \frac{2}{(1+at)^2} \left[1 - \frac{at}{1+at} \right]^{-3} - 1$$

$$= \frac{2}{(1+at)^2} (1+at)^3 - 1$$

$$E\{X^2(t)\} = 2(1+at) - 1 = 2 + 2at - 1 = \underline{\underline{1 + 2at}}$$

$$\boxed{E\{X^2(t)\} = 1 + 2at}$$

$$\text{Var}[X(t)] = E\{X^2(t)\} - [E(X(t))]^2 = 1 + 2at - 1 = 2at$$

$$\boxed{\text{Var}[X(t)] = 2at}$$

Ridhi Sethi

Since $E\{X^2(t)\}$ is not a constant; it is
function of t

And since $\text{Var}\{X(t)\}$ is not a constant, it is
function of t

\therefore Given process is not stationary

Ridhi Seth, Asst. Professor

Random Processes

- A random variable is a fn which assigns a real no. to every outcome of a random exp. while a random process is a fn that assigns a time fn to every outcome of a random exp.
- Set of fns $\{x_1(t), x_2(t), \dots, x_n(t)\}$ represents a random process.

Classification :

- 1) Discrete Random Sequence
- 2) Discrete Random process
- 3) Continuous Random Sequence
- 4) Continuous Random process
- 5) Deterministic Random process
- 6) Non-deterministic Random process

Statistical Averages of Random processes :

1) Mean: $\mu(t) = E[x(t)] = \int_{-\infty}^{\infty} x(t) \cdot f(x, t) dx$

- 2) Auto co-relation: Auto-correlation of a process $\{x(t)\}$ is denoted by $R_{xx}(t_1, t_2)$, $R_x(t_1, t_2)$ or $R(t_1, t_2)$

$$R_{xx}(t_1, t_2) = E[\{x(t_1)\} \{x(t_2)\}]$$

$$x(t) = A \sin \lambda t$$

- 3) Auto-covariance: auto co-variance of a process $\{x(t)\}$ is denoted by $C_{xx}(t_1, t_2)$ or $C_x(t_1, t_2)$ or $C(t_1, t_2)$.

$$C_{xx}(t_1, t_2) = E[\{x(t_1) - \mu(t_1)\} \{x(t_2) - \mu(t_2)\}]$$

$$C_{xx}(t_1, t_2) = E[\{x(t_1)\} \{x(t_2)\}] - \mu(t_1) \mu(t_2)$$

- 4) Co-relation coefficient:

$$\rho(t_1, t_2) = \frac{C_{xx}(t_1, t_2)}{\sqrt{C(t_1, t_1) C(t_2, t_2)}}$$

$$C(t_1, t_1) = \text{Variance of } x(t_1)$$

$$C(t_2, t_2) = \text{Variance of } x(t_2)$$

5) Cross co-relation: of $\{x(t)\}$ and $\{y(t)\}$ is denoted by $R_{xy}(t_1, t_2)$

$$R_{xy}(t_1, t_2) = E[\{x(t_1)\}\{y(t_2)\}]$$

6) Cross co-variance: of $\{x(t)\}$ and $\{y(t)\}$ is denoted by $C_{xy}(t_1, t_2)$

$$C_{xy}(t_1, t_2) = R_{xy}(t_1, t_2) - \mu_x(t_1)\mu_y(t_2)$$

7) Cross - co-relation coefficient:

$$\rho_{xy}(t_1, t_2) = \frac{C_{xy}(t_1, t_2)}{\sqrt{C_x(t_1, t_1) C_y(t_2, t_2)}}$$

$C_x(t_1, t_1) = \text{variance of } x(t_1)$

$C_y(t_2, t_2) = \text{variance of } y(t_2)$

Stationarity / Stationary Random Process: A random process is called a strongly stationary process or strict sense stationary (SSS) if all its distributions are invariant under translation of time.

First order Stationary Random process: A RP is said to be first order stationary if its density func. is invariant under the translation in time i.e.

$$f(x, t) = f(x, t + h)$$

Second Order Stationary random process: A RP is said to be second order stationary if the joint pdf of $\{x(t_1)\}$ and $\{x(t_2)\}$ are invariant under translation i.e.

$$f(x_1, x_2, t_1, t_2) = f(x_1, x_2, t_1 + h, t_2 + h)$$

Wide Sense Stationarity (WSS): A RP $\{x(t)\}$ is called a wide sense stationary process if

(i) $E[\{x(t)\}] = \text{constant}$

(ii) Auto co-relation depends only on time difference $E[x(t)x(t-\tau)] = R(\tau)$

$R_{xx}(t_1, t_2)$ is a fn of time difference $(t_1 - t_2)$ 105

Note:

- 1) In strict sense, SP is a WSS process but WSS process need not be a SSS process.
- 2) A RP that is not stationary in any sense whether in strict sense or in wide sense.
- 3) Two RP $\{x(t)\}$ $\{y(t)\}$ are said to be jointly stationary in wide sense if each process is individually a WSS process and cross co-relation $R_{xy}(t_1, t_2)$ is a fn of $t_1 - t_2$ only.
- 8) Show that RP $x(t) = A \cos(\omega_0 t + \theta)$ is WSS if A and ω_0 are constants and θ is uniformly distributed in the interval $(0, 2\pi)$

Sol For a process to be WSS

(i) $E[x(t)] = \text{const.}$

(ii) $E[x(t)x(t-\tau)] = R(\tau)$

$R(t_1, t_2)$ is a fn of $(t_1 - t_2)$ given is uniformly distributed R.V $f(\theta) = \begin{cases} 1/b-a & ; a < \theta < b \\ 0 & ; \text{o/w} \end{cases}$

$$f(\theta) = \frac{1}{2\pi}, 0 < \theta < 2\pi.$$

(i) $E[x(t)] = \text{const.}$

$$E[x(t)] = \int_0^{2\pi} x(t) f(\theta) d\theta.$$

$$= \frac{1}{2\pi} \int_0^{2\pi} A \cos(\omega_0 t + \theta) d\theta.$$

$$= \frac{A}{2\pi} \left[\sin(\omega_0 t + \theta) \right]_0^{2\pi}$$

$$= \frac{A}{2\pi} [\sin(2\pi + \omega_0 t) - \sin \omega_0 t]$$

$$= \frac{A}{2\pi} [\sin \omega_0 t - \sin \omega_0 t]$$

$$= 0(\text{const})$$

$$(ii) R_{XX}(t_1, t_2) = E[X(t_1) X(t_2)]$$

$$106 \quad X(t_1) = A \cos(\omega_0 t_1 + \theta)$$

$$X(t_2) = A \cos(\omega_0 t_2 + \theta)$$

$$R_{XX}(t_1, t_2) = E[A \cos(\omega_0 t_1 + \theta) \cdot A \cos(\omega_0 t_2 + \theta)]$$

$$= E\left[\frac{A^2}{2} 2 \cos(\omega_0 t_1 + \theta) \times \cos(\omega_0 t_2 + \theta)\right]$$

$$= E\left[\frac{A^2}{2} [\cos(\omega_0(t_1 + t_2) + 2\theta) + \cos(\omega_0(t_1 - t_2))]\right]$$

$$= \frac{A^2}{2} E[\cos(\omega_0(t_1 + t_2) + 2\theta) + \cos(\omega_0(t_1 - t_2))]$$

$$E[X + Y] = E[X] + E[Y]$$

$$R_{XX}(t_1, t_2) = \frac{A^2}{2} E[\cos(\omega_0(t_1 + t_2) + 2\theta)] + \frac{A^2}{2} E[\cos(\omega_0(t_1 - t_2))]$$

$$= \frac{A^2}{2} \int_0^{2\pi} \cos[\omega_0(t_1 + t_2) + 2\theta] f(\theta) d\theta + \frac{A^2}{2} \int_0^{2\pi} \cos(\omega_0(t_1 - t_2)) f(\theta) d\theta$$

$$= \frac{A^2}{4\pi} \left[\sin(\omega_0(t_1 + t_2) + 2\theta) \right]_0^{2\pi} + \frac{A^2}{4\pi} \left[\theta \cos \omega_0(t_1 - t_2) \right]_0^{2\pi}$$

$$R_{XX}(t_1, t_2) = \frac{A^2}{4\pi} \left[\sin[4\pi + \omega_0(t_1 + t_2)] - \sin \omega_0(t_1 + t_2) \right] +$$

$$\frac{A^2}{4\pi} \cos \omega_0(t_1 - t_2) [2\pi - 0]$$

$$= \frac{A^2}{4\pi} \left[\sin(\omega_0(t_1 + t_2)) - \sin \omega_0(t_1 + t_2) \right] + \frac{A^2}{2} \cos \omega_0(t_1 - t_2)$$

$$= \frac{A^2}{2} \cos \omega_0(t_1 - t_2)$$

Q) A. R. P. is defined as $X(t) = A \sin(\omega_0 t + \theta)$ where A and ω_0 are const if θ is uniformly distributed R.V in the interval $(-\pi, \pi)$. Show that $X(t)$ is WSS

Q) Show that the process $\{x(t)\} = A \cos \lambda t + B \sin \lambda t$, where A and B are R.V is a WSS process if (i) $E(A) = E(B) = 0$
 (ii) $E(A^2) = E(B^2)$ (iii) $E(AB) = 0$. 107

Sol For a process to be WSS $E[\{x(t)\}] = 0$.

$$\begin{aligned} \text{(a)} \quad E[x(t)] &= E[A \cos \lambda t + B \sin \lambda t] \\ &= E[A \cos \lambda t] + E[B \sin \lambda t] \\ &= \cos \lambda t E[A] + \sin \lambda t E[B] \quad \text{--- (1)} \end{aligned}$$

For WSS process if $E(A) = E(B) = 0$ then $E[x(t)] = 0$ in (1)
 Hence $E(A) = E(B) = 0$

$R_{xx}(t_1, t_2)$ must be fn of $(t_1 - t_2)$

$$\begin{aligned} R_{xx}(t_1, t_2) &= E[\{x(t_1)\} \{x(t_2)\}] \\ x(t_1) &= A \cos \lambda t_1 + B \sin \lambda t_1 \\ x(t_2) &= A \cos \lambda t_2 + B \sin \lambda t_2 \end{aligned}$$

$$\begin{aligned} R_{xx}(t_1, t_2) &= E[(A \cos \lambda t_1 + B \sin \lambda t_1)(A \cos \lambda t_2 + B \sin \lambda t_2)] \\ &= E[A^2 \cos \lambda t_1 \cos \lambda t_2 + B^2 \sin \lambda t_1 \sin \lambda t_2 + \\ &\quad AB \sin(\lambda t_1 + \lambda t_2)] \\ &= \cos \lambda t_1 \cos \lambda t_2 E(A^2) + \sin \lambda t_1 \sin \lambda t_2 E(B^2) + \\ &\quad \sin(\lambda t_1 + \lambda t_2) E[AB] \end{aligned}$$

$$\begin{aligned} \text{if } E(A^2) &= E(B^2) = k \\ E[AB] &= 0. \end{aligned}$$

$$\begin{aligned} R_{xx}(t_1, t_2) &= k [\cos \lambda t_1 \cos \lambda t_2 + \sin \lambda t_1 \sin \lambda t_2] \\ &= k \cos \lambda(t_1 - t_2). \end{aligned}$$

Auto. co-relation fn is func of $(t_1 - t_2)$

Q) Two R.P's $X(t)$ and $Y(t)$ are given by $X(t) = A \cos \omega_0 t + B \sin \omega_0 t$, $Y = B \cos \omega_0 t - A \sin \omega_0 t$. Show that $X(t)$ and $Y(t)$ are jointly wide. Show that A, B are uncorrelated R.V with zero mean and same variances and ω_0 is a const.

$$\Delta 10 \quad \text{cov}(A, B) = E[AB] - E[A]E[B]$$

$$E(A^2) = E(B^2)$$

$$108 \quad \mu_A = E(A) = 0$$

$$\mu_B = E(B) = 0$$

$$\text{cov}(A, B) = 0$$

$$0 = E[AB] - 0$$

$$E[AB] = 0$$

$$\text{Var}(A) = E[A^2] - [E(A)]^2$$

$$\text{Var}(A) = E(A^2)$$

$$\text{Var}(B) = E(B^2)$$

$$E(A^2) = E(B^2)$$

$$(i) \quad E[x(t)] = E[A \cos \omega_0 t + B \sin \omega_0 t]$$

$$E[x(t)] = E[A \cos \omega_0 t] + B E[B \sin \omega_0 t]$$

$$= \cos \omega_0 t E[A] + \sin \omega_0 t E[B]$$

$$\because E[A] = E[B] = 0$$

$$E[x(t)] = 0$$

$$(ii) \quad R_{xx}(t_1, t_2) = E[\{x(t_1)\} \{x(t_2)\}]$$

$$= E[(A \cos \omega_0 t_1 + B \sin \omega_0 t_1)(A \cos \omega_0 t_2 + B \sin \omega_0 t_2)]$$

$$= E[A^2 \cos \omega_0 t_1 \cos \omega_0 t_2 + B^2 \sin \omega_0 t_1 \sin \omega_0 t_2 +$$

$$AB \cos \omega_0 t_1 \sin \omega_0 t_2 + AB \sin \omega_0 t_1 \cos \omega_0 t_2]$$

$$= \cos \omega_0 t_1 \cos \omega_0 t_2 E[A^2] + \sin \omega_0 t_1 \sin \omega_0 t_2 E[B^2]$$

$$= k [\cos \omega_0 t_1 \cos \omega_0 t_2 + \sin \omega_0 t_1 \sin \omega_0 t_2]$$

$$= k \cos \omega_0(t_1 - t_2)$$

$\{x(t)\}$ is WSS

& $y(t)$ is WSS

$$(iii) R_{xy}(t_1, t_2) = E\{x(t_1)x(t_2)\}$$

109

$$\{x(t_1)\} = A \cos \omega_0 t_1 + B \sin \omega_0 t_1$$

$$\{y(t_2)\} = B \cos \omega_0 t_2 - A \sin \omega_0 t_2$$

$$\begin{aligned} R_{xy}(t_1, t_2) &= E\{[A \cos \omega_0 t_1 + B \sin \omega_0 t_1][B \cos \omega_0 t_2 - A \sin \omega_0 t_2]\} \\ &= E\{AB \cos \omega_0 t_1 \cos \omega_0 t_2 - A^2 \cos \omega_0 t_1 \sin \omega_0 t_2 \\ &\quad + B^2 \sin \omega_0 t_1 \cos \omega_0 t_2 - AB \sin \omega_0 t_1 \sin \omega_0 t_2\} \end{aligned}$$

$$\begin{aligned} &= \cos \omega_0 t_1 \cos \omega_0 t_2 E\{AB\} - \cos \omega_0 t_1 \sin \omega_0 t_2 E\{A^2\} \\ &\quad + \sin \omega_0 t_1 \cos \omega_0 t_2 E\{B^2\} - \sin \omega_0 t_1 \sin \omega_0 t_2 E\{AB\} \end{aligned}$$

$$E\{AB\} = 0 \quad E\{A^2\} = E\{B^2\} = k$$

$$\begin{aligned} R_{xy}(t_1, t_2) &= k[\sin \omega_0 t_1 \cos \omega_0 t_2 - \cos \omega_0 t_1 \sin \omega_0 t_2] \\ &= k[\sin(\omega_0 t_1 - \omega_0 t_2)] \\ &= k \sin \omega_0(t_1 - t_2) \end{aligned}$$

Ergodicity:

Time Average: If $\{x(t)\}$ is the R.P. Then $\frac{1}{2T} \int_{-T}^T x(t) dt$ is called the time avg of $\{x(t)\}$ over the interval $-T$ to T . It is denoted by μ .

Ergodic Process: A R.P. is said to be ergodic if its statistical averages (ensemble) are equal to time averages.

Mean ergodic process: If a R.P. $\{x(t)\}$ has a const. mean, $E\{x(t)\} = \mu$ and $\lim_{T \rightarrow \infty} \bar{x}_T = \mu$.

Co-relation ergodic process: The stationary process $\{x(t)\}$ is said to be co-relation ergodic if the process.

$Z(t) = \{x(t+\lambda)x(t)\}$ is mean ergodic and if

$$\lim_{T \rightarrow \infty} \bar{Z}_T = R(\lambda)$$

$$\text{where } \bar{Z}_T = \frac{1}{2T} \int_{-T}^T x(t+\lambda)x(t) dt.$$

$$R_{xx}(\lambda) = E[X(t+\lambda)X(t)]$$

$$110 \quad R_{xx}(t_1, t_2) = E[X(t_1)X(t_2)]$$

$$Z_T = X(t+\lambda)X(t)$$

Q) If WSS process $\{X(t)\}$ is given by $10 \cos(100t + \theta)$ where θ is uniformly distributed over the interval $(-\pi, \pi)$. Prove that $\{X(t)\}$ is co-relation ergodic.

$$\text{Sol} \quad \lim_{T \rightarrow \infty} \frac{1}{T} \bar{Z}_T = R(\lambda)$$

$$R(\lambda) = E[X(t+\lambda)X(t)]$$

$$X(t+\lambda) = 10 \cos[100(t+\lambda) + \theta]$$

$$\begin{aligned} R(\lambda) &= E[10 \cos\{100(t+\lambda) + \theta\} 10 \cos(100t + \theta)] \\ &= E\left[\frac{100}{2} \times 2 \cos(100t + 100\lambda + \theta) \cos(100t + \theta)\right] \\ &= E[50 \cos(200t + 200\lambda + 2\theta) + 50 \cos(100\lambda)] \\ &= 50 E[\cos(200t + 200\lambda + 2\theta)] + 50 E[\cos(100\lambda)] \end{aligned}$$

$$f(\theta) = \frac{1}{b-a}; \quad a < \theta < b$$

$$f(\theta) = \frac{1}{2\pi}; \quad -\pi < \theta < \pi$$

$$\begin{aligned} R(\lambda) &= 50 \int_{-\pi}^{\pi} \frac{1}{2\pi} \cos(200t + 100\lambda + 2\theta) d\theta + 50 \int_{-\pi}^{\pi} \cos(100\lambda) \left(\frac{1}{2\pi} d\theta\right) \\ &= \frac{50}{2\pi} [\sin(2\pi + 200t + 100\lambda) - \sin(-2\pi + 200t + 100\lambda)] \\ &\quad + \frac{50}{2\pi} \cos(100\lambda)(2\pi) \end{aligned}$$

$$\sin(-x) = -\sin x$$

$$\sin(2\pi - x) = -\sin x$$

$$\sin(2\pi + x) = \sin x$$

$$\begin{aligned} R(\lambda) &= \frac{25}{\pi} [\sin(200t + 100\lambda) - \sin(200t + 100\lambda)] + 50 \cos(100\lambda) \\ R(\lambda) &= 50 \cos(100\lambda) \end{aligned}$$

$$\bar{Z}_T = \frac{1}{2T} \int_{-T}^T \lambda(t+\lambda) x(t) dt$$

|||

$$= \frac{1}{2T} \int_{-T}^T 10 \cos(100t + 100\lambda + \theta) 10 \cos(100t + \theta) dt$$

$$= \frac{1}{2T} \int_{-T}^T 50 [\cos(200t + 100\lambda + 2\theta) + \cos(100\lambda)] dt$$

$$= \frac{25}{T} \int_{-T}^T \cos(200t + 100\lambda + 2\theta) dt + \frac{25}{T} \cos(100\lambda) \int_{-T}^T dt$$

$$= \frac{25}{T} \int_{-T}^T \cos(200t + 100\lambda + 2\theta) + 50 \cos(100\lambda)$$

$$= 0 + 50 \cos(100\lambda)$$

$$= R(\lambda)$$

Q) The process $\{x(t)\}$ whose probability distribution under certain cond. is given by

$$P[x(t)] = \frac{(at)^{n-1}}{(1+at)^{n+1}} \quad n = 1, 2, 3, \dots$$

$$= \frac{at}{1+at} \quad ; n=0$$

Sol Process to be SSS

(i) $E[x(t)]$ should be const.

(ii) $E[x^2(t)]$ should be const

(iii) $\text{Var}[x(t)]$ should be const.

(iv) $R_{xx}(t_1, t_2)$ should be fn $(t_1 - t_2)$

x_i	0	1	2	3	4	...
$P[x(t)]$	$\frac{at}{1+at}$	$\frac{a^2}{(1+at)^2}$	$\frac{at}{(1+at)^3}$	$\frac{(at)^2}{(1+at)^4}$
$n=1, 2, 3, \dots$						

$$(i) E[x(t)] = \sum_{n=0}^{\infty} n P[x(t)]$$

$$= \left(0 \times \frac{at}{1+at} \right) + \frac{1}{(1+at)} + \frac{2at}{(1+at)^3} + \frac{3(at)^2}{(1+at)^4} + \dots \infty$$

$$= \frac{1}{(1+at)^2} \left[1 + \frac{2at}{(1+at)} + \frac{3(at)^2}{(1+at)^2} + \dots \infty \right]$$

112

$$\text{let } \frac{at}{1+at} = x.$$

$$\begin{aligned} E[x(t)] &= \frac{1}{(1+at)^2} [1 + 2x + 3x^2 + 4x^3 + \dots \infty] \\ &= \frac{1}{(1+at)^2} [1-x]^{-2} \\ &= \frac{1}{(1+at)^2} \left[1 - \frac{at}{1+at} \right]^{-2} \\ &= \frac{1}{(1+at)^2} \left[\frac{1+at-at}{1+at} \right]^{-2} \\ &= \frac{1}{(1+at)^2} \left(\frac{1}{1+at} \right)^{-2} \\ &= 1 \end{aligned}$$

$$(ii) E[x^2(t)] = \sum_{n=0}^{\infty} n^2 P[x(t)]$$

$$= \sum_{n=0}^{\infty} n(n+1) P[x(t)] - \sum_{n=0}^{\infty} n P[x(t)]$$

Replace n^2 with $(n(n+1) - n)$

$$= \sum_{n=0}^{\infty} [n(n+1) - n] P[x(t)]$$

$$= \sum_{n=0}^{\infty} [n(n+1) P[x(t)] - n P[x(t)]]$$

$$= \left[\frac{2}{(1+at)^2} + \frac{2(3at)}{(1+at)^3} + \frac{3 \times 4(at)^2}{(1+at)^4} + \dots \infty \right] - E[x(t)]$$

$$= \frac{2}{(1+at)^2} \left[1 + \frac{3at}{(1+at)} + 6 \left[\frac{at}{1+at} \right]^2 + \dots \infty \right] - 1$$

$$= \frac{2}{(1+at)^2} [1 + 3x + 6x^2 + 10x^3 + \dots \infty] - 1$$

$$= \frac{2}{(1+at)^2} (1-x)^{-3} - 1$$

$$= \frac{2}{(1+at)^2} \left[1 - \frac{at}{1+at} \right]^{-3} - 1$$

$$= \frac{2}{(1+at)^2} (1+at)^3 - 1$$

$$= 2(1+at) - 1$$

$$= 2at + 1$$

$E[x^2(t)]$ is not a const

113

Hence, this process is not a SSS process

$$\begin{aligned}\text{(ii) } \text{Var}[x(t)] &= E[x^2(t)] - E[x(t)]^2 \\ &= 1 + 2at - 1 \\ &= 2a \cdot t\end{aligned}$$

This process is not a stationary process.

Power spectral density function: If $\{x(t)\}$ is a stationary process either in strict sense or in wide sense with auto co-relation fn given as $R(\tau)$ then the fourier transform of $R(\tau)$ is called the power spectral density fn of $\{x(t)\}$ denoted by $S_{xx}(\omega)$ or $S_x(\omega)$ or $S(\omega)$

$$S(\omega) = \int_{-\infty}^{\infty} R(\tau) e^{i\omega\tau} d\tau \quad \text{--- (1)}$$

$$S(f) = \int_{-\infty}^{\infty} R(\tau) e^{-i2\pi f\tau} d\tau \quad \text{--- (2)}$$

Weiner Khinchins Relation: The auto correlation $R(\tau)$ is given by fourier inverse transform $S(\omega)$

$$R(\tau) = \frac{1}{2\pi} \int_{-\infty}^{\infty} S(\omega) e^{i\omega\tau} d\omega \quad \text{--- (3)}$$

$$R(\tau) = \int_{-\infty}^{\infty} S(f) e^{i2\pi f\tau} df \quad \text{--- (4)}$$

- If $x(t)$ and $y(t)$ are 2 jointly stationary processes with cross co-relation fn $R_{xy}(\tau)$. Then fourier transform of $R_{xy}(\tau)$ is called cross power spectral density of $\{x(t)\}$ and $\{y(t)\}$.

Properties of power spectral density fn:

1) The value of spectral density fn at zero frequency is equal to the total area under the graph of the auto-co-relation fn.

$$S(0) = \int_{-\infty}^{\infty} R(\tau) d\tau \quad \text{--- (5)}$$

$$S(0) = \int_{-\infty}^{\infty} R(\tau) d\tau \quad \text{--- (6)}$$

- 2) The mean square value of a WSS process is equal to the total area under the graph of spectral density

114

$$E[X^2(t)] = R(0)$$

put $T=0$ in (4)

$$R(0) = \int_{-\infty}^{\infty} S(f) df.$$

- 3) The spectral density fn of a real R.P is an even fn
 $S(-\omega) = S(\omega)$

- 4) The spectral density & auto co-relation fn of a real WSS process form a fourier cosine transform pair

Proof: $\mathcal{F} S(\omega) = \int_{-\infty}^{\infty} R(\tau) e^{-i\omega\tau} d\tau$

$$= \int_{-\infty}^{\infty} R(\tau) (\cos \omega\tau - i \sin \omega\tau) d\tau$$

$$= \int_{-\infty}^{\infty} R(\tau) \cos \omega\tau d\tau - i \underbrace{\int_{-\infty}^{\infty} R(\tau) \sin \omega\tau d\tau}_{\text{odd}}$$

$$= 2 \int_0^{\infty} R(\tau) \cos \omega\tau d\tau - 0.$$

$$= \int_0^{\infty} 2 R(\tau) \cos \omega\tau d\tau.$$

$S(\omega)$ = fourier cosine transform of $2R(\tau)$

$$R(\tau) = \frac{1}{2\pi} \int_{-\infty}^{\infty} S(\omega) e^{i\omega\tau} d\omega.$$

$$= \frac{1}{2\pi} \int_{-\infty}^{\infty} S(\omega) (\cos \omega\tau + i \sin \omega\tau) d\omega.$$

$$= \frac{1}{2\pi} \int_{-\infty}^{\infty} S(\omega) \cos \omega\tau d\omega + \frac{i}{2\pi} \int_{-\infty}^{\infty} S(\omega) \sin \omega\tau d\omega$$

$$= \frac{1}{\pi} \int_0^{\infty} S(\omega) \cos \omega\tau d\omega$$

3) Wiener Khinchine Theorem: If $X_T(\omega)$ is the Fourier transform of RP defined as

115

$$X_T(t) = \begin{cases} x(t) & ; |t| \leq T \\ 0 & ; |t| > T \end{cases}$$

where $\{X_T(t)\}$ is a real WSS process.

Q) Find the power spectral density of a random telegraph signal with auto-correlation fn of the telegraph signal given by $R(\tau) = -a^2 e^{-2\gamma|\tau|}$

Sol.

$$\begin{aligned} S(\omega) &= \int_{-\infty}^{\infty} R(\tau) e^{-i\omega\tau} d\tau \\ &= \int_{-\infty}^{\infty} a^2 e^{-2\gamma|\tau|} e^{-i\omega\tau} d\tau \\ &= \int_{-\infty}^0 a^2 e^{2\gamma\tau} e^{-i\omega\tau} d\tau + \int_0^{\infty} a^2 e^{-2\gamma\tau} e^{-i\omega\tau} d\tau \\ &= a^2 \int_{-\infty}^0 e^{\tau(2\gamma - i\omega)} d\tau + a^2 \int_0^{\infty} e^{-\tau(2\gamma + i\omega)} d\tau \\ &= a^2 \left[\frac{e^{-\tau(2\gamma - i\omega)}}{2\gamma - i\omega} \right]_{-\infty}^0 + a^2 \left[\frac{e^{-\tau(2\gamma + i\omega)}}{2\gamma + i\omega} \right]_0^{\infty} \\ &= a^2 \left[\frac{1}{2\gamma - i\omega} \right] + a^2 \left[\frac{-1}{2\gamma + i\omega} \right] \\ &= \frac{a^2}{2\gamma - i\omega} + \frac{a^2}{2\gamma + i\omega} = \frac{4\gamma a^2}{4\gamma^2 + \omega^2} \end{aligned}$$

Q) Find the power spectral density of a WSS process with auto co-relation fn given as $R(\tau) = e^{-\alpha\tau^2}$

Sol.

$$\begin{aligned} S(\omega) &= \int_{-\infty}^{\infty} R(\tau) e^{-i\omega\tau} d\tau \\ &= \int_{-\infty}^{\infty} e^{-\alpha\tau^2} e^{-i\omega\tau} d\tau \\ &= \int_{-\infty}^{\infty} e^{(-\alpha\tau^2 - i\omega\tau)} d\tau \\ &= \int_{-\infty}^{\infty} e^{-\alpha\left(\tau^2 + \frac{i\omega\tau}{\alpha}\right)} d\tau \end{aligned}$$

add & sub $\frac{\omega^2}{4\alpha^2}$ to power of e

$$S(\omega) = \int_{-\infty}^{\infty} e^{-\alpha\left(\tau^2 + \frac{i\omega\tau}{\alpha} - \frac{\omega^2}{4\alpha^2} + \frac{\omega^2}{4\alpha^2}\right)} d\tau$$

$$\begin{aligned}
 11b \quad &= \int_{-\infty}^{\infty} e^{-\alpha \left(\tau^2 + \frac{i\omega\tau}{\alpha} + \frac{\omega^2}{4\alpha^2} \right) - \alpha \left(\frac{\omega^2}{4\alpha^2} \right)} d\tau \\
 &= \int_{-\infty}^{\infty} e^{-\alpha \left(\tau + \frac{i\omega}{2\alpha} \right)^2} e^{-\alpha \left(\frac{\omega^2}{4\alpha^2} \right)} d\tau \\
 &= e^{-\omega^2/4\alpha} \int_{-\infty}^{\infty} e^{-\alpha \left(\tau + \frac{i\omega}{2\alpha} \right)^2} d\tau
 \end{aligned}$$

$$\text{put } \sqrt{\alpha} \left(\tau + \frac{i\omega}{2\alpha} \right) = t$$

$$\sqrt{\alpha} d\tau = dt$$

$$d\tau = \frac{dt}{\sqrt{\alpha}}$$

$$S(\omega) = e^{-\omega^2/4\alpha} \int_{-\infty}^{\infty} e^{-t^2} \frac{dt}{\sqrt{\alpha}}$$

$$e^{-t^2} = \text{even fn.}$$

$$= \frac{e^{-\omega^2/4\alpha}}{\sqrt{\alpha}} \cdot 2 \int_0^{\infty} e^{-t^2} dt$$

$$\text{put } t^2 = k$$

$$2t dt = dk$$

$$S(\omega) = \frac{2 e^{-\omega^2/4\alpha}}{\sqrt{\alpha}} \int_0^{\infty} e^{-k} \frac{dk}{2\sqrt{k}}$$

$$= \frac{e^{-\omega^2/4\alpha}}{\sqrt{\alpha}} \int_0^{\infty} e^{-k} k^{1/2-1} dk$$

$$= \frac{e^{-\omega^2/4\alpha}}{\sqrt{\alpha}} \sqrt{1/2} = \frac{e^{-\omega^2/4\alpha}}{\sqrt{\alpha}} \sqrt{\pi}$$

$$= \sqrt{\frac{\pi}{\alpha}} e^{-\omega^2/4\alpha}$$